DOCUMENT RESUME

ED 144 762 RC 010 115

AUTHOR Sharda, Bam Dev; Elder, Joseph W.
TITLE Industrialization and the Processes of Stratification in Rural Societies: A Comparison of Rural India and Rural United States.
PUB DATE 25 Aug 75
NOTE 32p.; Paper presented at the Annual Meeting of the American Sociological Association (70th, August 25-29, 1975)

EDRS PRICE MF-$0.83 HC-$2.06 Plus Postage.
DESCRIPTORS Age Differences; *Comparative Analysis; Foreign Countries; Males; *Models; *Occupations; Parental Background; *Rural Urban Differences; *Status
IDENTIFIERS Blau (Peter M); Cross National Studies; Duncan (Otis Dudley); *India; Standard International Occupational Prestige Scale; *United States

ABSTRACT

Status attainment processes of the experienced civilian labor force of rural India and rural United States were analyzed by using the Blau-Duncan model, for the first time employing that model to compare rural sectors of an agrarian society with those in an industrial society. United States data were obtained from the 1962 Occupational Changes in a Generation Survey; Indian data were obtained from the 1962 Census of India Village Survey studies of three states. The comparison was limited to male respondents 20 to 64 years old. Treiman's Standard International Occupational Prestige Scale was used to measure occupational status. Because of lack of data from the India sample, the variable of father's education was eliminated from the basic Blau-Duncan model. Using path analysis, results were obtained for the total sample as well as for different age cohorts. Results showed education was relatively unimportant in the rural Indian occupational structure but relatively important in the US structure. The largest direct effect on current occupation in both societies was from the respondent's previous occupation (India) or first job (US). Effect of the third variable, father's occupation, on current occupation was more important in India than in the US. (RS)

****
* Documents acquired by ERIC include many informal unpublished *
* materials not available from other sources. ERIC makes every effort *
* to obtain the best copy available. Nevertheless, items of marginal *
* reproducibility are often encountered and this affects the quality *
* of the microfiche and hardcopy reproductions ERIC makes available *
* via the ERIC Document Reproduction Service (EDRS). EDRS is not *
* responsible for the quality of the original document. Reproductions *
* supplied by EDRS are the best that can be made from the original. *
****

ED144762

Session 19-3
Monday
August 25
10:30 a.m.

INDUSTRIALIZATION AND THE PROCESSES OF STRATIFICATION IN RURAL SOCIETIES: A COMPARISON OF RURAL INDIA AND RURAL UNITED STATES

Bam Dev Sharda
University of Utah

and

Joseph W. Elder
University of Wisconsin

U S DEPARTMENT OF HEALTH, EDUCATION & WELFARE
NATIONAL INSTITUTE OF EDUCATION

THIS DOCUMENT HAS BEEN REPRODUCED EXACTLY AS RECEIVED FROM THE PERSON OR ORGANIZATION ORIGINATING IT. POINTS OF VIEW OR OPINIONS STATED DO NOT NECESSARILY REPRESENT OFFICIAL NATIONAL INSTITUTE OF EDUCATION POSITION OR POLICY

A paper prepared for presentation at the 70th Annual Meeting of the American Sociological Association, August 25-29, 1975.

RC010115

Not for reproduction, quotation, or any other use without permission of the authors, except for press use after 10:30 a.m., August 25.

"PERMISSION TO REPRODUCE THIS MATERIAL HAS BEEN GRANTED BY
Bam Dev Sharda

TO THE EDUCATIONAL RESOURCES INFORMATION CENTER (ERIC) AND THE ERIC SYSTEM CONTRACTORS."

Permission Granted to ERIC
19-01-76

INDUSTRIALIZATION AND THE PROCESSES OF STRATIFICATION IN RURAL SOCIETIES: A COMPARISON OF RURAL INDIA AND RURAL UNITED STATES

Bam Dev Sharda
University of Utah

Joseph W. Elder
University of Wisconsin

This paper examines the status attainment processes of the "Experienced Civilian Labor Force" in two different societies: rural India and rural United States. The Census definition of "rural" is employed in both cases. The U.S. data were obtained from Occupational Changes in a Generation Survey (1962). The Indian data were obtained from the Census of India's Village Survey studies of three states (1962). The analysis was restricted to 20-64 year-old male respondents in 1962. Treiman's STANDARD INTERNATIONAL OCCUPATIONAL PRESTIGE SCALE was used to measure occupational status in both cases. The basic Blau and Duncan model was modified by dropping father's education from the model to conform to our India data. By using path analysis, the results were obtained for the total sample and for different age cohorts. Our tentative conclusions are as follows:

1.[a] Inheritance defined as $r^2_{fs(India)} = .676^2 = .47$ $r^2_{fs(U.S.)} = .039^2 = 0$

2.[b] Direct Effects of Education: $Py_4v = .115$(India) $Py_4v = .266$(U.S.)

3.[c] Direct Effects of Father's Occupation:

$Py_1x = .137$ to $Py_4x = .49$ (rises with age in India)

$Py_1x = .060$ to $Py_4x = -.012$ (falls with age, if anything, in the U.S.)

4. Total impact of all variables on Y_4:

Age	India R^2	U.S.
25-34	.87	.25
35-44	.74	.15
45-54	.67	.15
55-64	.55	.11

5. As age increases in India, previous occupation has a diminishing effect, from large to moderate (.837, .673, .493, .370); as age increases in the U.S., first job status also has a diminishing effect, from moderate to small (.296, .124, .147, .148).

These differences are attributed to the differences in the levels of industrialization of these two rural societies.

---

[a] f = father's occupational status at death or retirement in the case of rural India and at son's age 16 in the case of rural U.S. s = current occupational status for age cohort 55-64.

[b] y_4 = occupational status at age cohort 55-64. v = respondent's education.

[c] y = current occupational status (y_1 at age cohort 25-34; y_4 at age cohort 55-64). x = father's occupational status.

INDUSTRIALIZATION AND THE PROCESSES OF STRATIFICATION IN RURAL SOCIETIES: A COMPARISON OF RURAL INDIA AND RURAL UNITED STATES

Bam Dev Sharda
University of Utah

Joseph W. Elder
University of Wisconsin

This paper compares the status attainment processes of two rural societies -- rural India and rural United States. We have followed the Blau and Duncan paradigm[1] of research in describing the process of stratification of these two societies. This paradigm reconceptualizes social mobility in terms of the life cycle of individuals[2]; taking origins, occupational entry, and occupational destination as major points in the life cycle (Blau and Duncan, 1967, especially chap. 5). Intersocietal comparisons using this model have been suggested (Treiman, 1970), and, in a few instances, this model has been applied to developing societies (Iutaka and Bock, 1973; Safar, 1971). However, no study so far has utilized this model either for analyzing mobility in agrarian societies, or for comparing the process of mobility in agrarian societies with that in industrial societies. We have, therefore, decided to focus our attention on this particular issue by comparing rural sectors of an industrial society (the United States) with those of an agrarian society (India).

A recent comparison of mobility data shows that the two most important variables indicating the "rates of mobility" in different countries are the "percentage of rural households" and the "degree of communication development" (Cutright, 1968). The fact that our samples are from only rural sectors reduces the relevance of the first variable, leaving intact the variable "degree

of communication development" and its correlates. "Degree of communication development," however, is frequently associated with the degree of industrialization, in fact, it is sometimes seen to be a consequence of industrialization (Lenski, 1966:301). Though our immediate focus in this study is to estimate the "analytic indicator" of a social process (stratification)[3] any differences we observe may be attributed to the two societies' differing degrees of industrialization.

Data

The data for our rural United States sample were obtained from the OCG-I survey of 1962.[4] The methods of sampling and data collection techniques are explained elsewhere (Blau and Duncan, 1967:10-19). The data for our rural Indian sample consists of 906 "heads of households" from 11 villages of the states of Punjab, Haryana and Himachal Pradesh. These 11 villages were selected purposely for this survey; they represent a deliberately broad range of cultural and dialectic-linguistic variations within the northwestern region of India. The 11 villages also differ in terms of their jati (intermarrying-caste) composition -- some are single jati villages while others are multi-jati villages. These villages also represent different levels of economic development and productive resources. These variations are reflected in the villages' different sizes and degrees of isolation from communication centers (roads and cities). These data were collected in 1962-63 as part of a large study of the Census of India.[5] Elsewhere we have suggested that this purposive sampling has probably produced a roughly representative sample of rural Indian male heads of households (Sharda, 1974:chap. 4). The term "rural" in this study is defined in accord with the Census classification of their respective countries.[6]

Unit of Analysis and the Variables

In order to conform to the definition of "experienced civilian male labor force," we restricted our Indian sample to ages 20 to 64. This resulted in the elimination of female heads of households, the older males beyond the age of 64, and a few younger males below the age of 20. All of the variables in our model were scaled in the common metric. For example Treiman's STANDARD INTERNATIONAL OCCUPATIONAL PRESTIGE SCALE scores (Treiman, forthcoming) were assigned to father's "current occupation" in both samples; and also to "previous occupation" in India and to its comparable variable, the respondent's "first job" in the United States. Elsewhere we have justified the rough comparability of "previous occupation" in India and "first job" in the United States (Sharda, 1974). Similarly we have recoded the variable of education in the OCG-I data to conform to the measure of education in approximate number of years in our Indian Sample.[7]

Modified Basic Model for the Process of Status Attainment

It is only recently that models of status attainment have been employed for describing the processes of stratification (Blau and Duncan, 1967:167). For the purposes of comparing the processes of stratification of rural India with those of rural United States, however, we have modified this basic model by eliminating the variable of father's education. The elimination of father's education as a predetermined variable from our model was necessitated by the absence of such data in our rural Indian sample. We hope that no serious distortions were introduced by this modification in our analysis. As a matter of fact, most developing societies exhibit a very limited range along this variable (see, for example, urban Brazil: Iutaka and Bock, 1973:218). The

rural U.S., despite its potentiality for a substantively large path in this direction, has simply been forced to conform to our requirements for comparison. We shall present an empirical estimation of this model prior to our discussion of the summary statistics.

Comparisons: Structure and Patterns of Status Transmission

Table 1 describes the association between the four selected variables of father's occupation, respondent's education, previous occupation in the case of rural India, and first job in the case of the rural U.S., and current occupation of the respondent. Examining the means and standard deviations,

Table 1 around here

we notice striking similarities within the occupational structures (in terms of their prestige dimensions) of the two countries.

On the average, the United States sample is about two points lower than the Indian sample in the origin status (father's occupation). Father's occupation exhibits a little more inequality in the U.S. (standard deviation 16.279) than in India (standard deviation 10.326). However, with regard to respondent's current occupation, the U.S. has a higher average ($\bar{X}$ = 39.715) than India ($\bar{X}$ = 37.987). Examination of the standard deviations with regard to origin and current occupation variables indicates that occupational inequality is more pronounced at the origin status in the U.S. (S_{faocc} = 16.279, S_{occ} = 13.645), and at the current occupation status in India (S_{occ} = 11.827, S_{faocc} = 10.326). The differences in the mean occupational prestige of previous occupation in India and first job in the U.S. simply may reflect age variations

at the point of respondent's entry into the occupational structure. It may also indicate that there is more intergenerational inheritance of agricultural occupations among rural Indian heads of households. Thus, sons of farmers in the U.S., rather than taking up their father's occupation, may accept lower status jobs and then continue to move up by changing jobs. The second interpretation seems the more likely; inasmuch as it is more consistent both with earlier notions about India and with a recent analysis of the OCG-I (U.S.) data using Treiman's SIOP scores (see Featherman et al., 1974b:Table 4c). The most pronounced differences between the two societies, however, concerns the variable of respondent's education. Respondents in our rural United States sample have considerably higher levels of education than rural Indians. The mean for the U.S. sample is about 10 years of education, whereas the corresponding mean for the rural Indian sample is about one year of education.

Turning to the comparison of zero-order correlations as a summary statistic of the relationship between similar variables within rural India and rural U.S., we find a number of interesting patterns. Beginning with the correlations between respondent's current occupation and the other three status variables, we discover that in both rural India and rural U.S. the highest correlation is between respondent's current occupation and respondent's previous occupation (in India) ($r_{occ,pocc} = .800$) or first job (in U.S.) ($r_{occ,fjob} = .310$). This suggests that in both countries there are important positive correlations of entry status with current occupation. But in India the correlation is considerably greater than in the United States.

Comparing the correlations between entry status and respondent's current occupation, on one hand, and father's occupation and respondent's current occupation, on the other, in both rural India and rural U.S. the correlation

with father's occupation, although marked (.673 in India; .120 in U.S.), is still lower than the correlation with entry status (.800 in India; .310 in the U.S.). In India the correlation between father's occupation and respondent's current occupation is considerably greater than in the U.S.

Comparing the correlations between education and respondent's current occupation, the picture becomes more complex. In rural U.S. the correlation between education and respondent's current occupation is relatively high (.303); in fact, the correlation is only slightly lower than that between entry status and respondent's current occupation (.310). In rural India the correlation between education and respondent's current occupation is practically zero (.069). This relative unimportance of education in the rural Indian occupational structure compared to the relative importance of education in the rural U.S. occupational structure suggests a striking difference between the rural sectors of the two societies.

Further evidence of the lack of importance of education in rural India can be seen by comparing the correlations between education and respondent's previous occupation (in India) and respondent's first job (in U.S.). The correlation is somewhat substantial in rural U.S. (.254); whereas it is practically zero in rural India (.039).

In neither rural India nor rural U.S. does father's occupation have a high correlation with respondent's education (.001 in India; .102 in U.S.). In both societies, father's occupation has a somewhat higher correlation with respondent's previous occupation (in India) and respondent's first job (in U.S.). But the correlation is markedly more substantial in rural India (.585) than it is in rural U.S. (.156).

In summary, according to our zero-order correlations, the societies appear similar to the extent that in both rural India and rural U.S. respondent's entry status -- previous occupation (in India) and respondent's first job (in U.S.) -- correlate relatively highly with respondent's current occupation. Then the differences between the societies become apparent. In India, the consistently higher correlations of status variables with father's occupation and lower correlations of status variables with respondent's education suggest rather sharply differing processes of status attainment in the two societies.

The analysis of these summary statistics thus suggests significant differences between rural India and rural United States -- even when the analysis also points out the similarities in the processes of status attainment in these two societies. We now shift to the multivariate analysis of the data in order to determine the relative weight of the background and career contingency factors in the status attainment processes of our two rural societies.

Paths of Attainment

The following path diagram estimates the path coefficients[8] of these relationships in both of these societies.

Figure 1 about here

The path coefficients expressed in Figure 1 are essentially the standardized regression coefficients. The results of the regression analysis for these three equations for both samples are reported in Table 2.

Table 2 about here

Raw regression coefficients are also computed along with the standardized regression coefficients. With common measures of the variables, intersocietal comparisons of this process may be better understood through the raw (unstandardized) regression coefficients. With raw-regressions there is an advantage of comparing the effects of mean differences among societies in terms of intercepts of the structural equations.

An examination of Table 2 and Figure 1 indicates that the largest direct effect on respondent's current occupation in both societies comes from the previous occupation (first job) of the respondent. The direct effect of father's occupation on current occupation remains sizeable in the case of India but has declined somewhat compared to the size of the zero-order correlation (Table 1) in the case of the U.S.

India and the U.S. differ widely in terms of the intercepts in the equations pertaining to each country where respondent's education is the dependent variable. (See Figure 2)

Figure 2 around here

Respondent's education cannot be predicted from father's occupational status at all in the case of rural India. Occupational status of father also seems to be a poor predictor of respondent's education in our U.S. sample.

In sum, status-attainment differences between rural India and rural U.S. are reflected in the fact that in India the major direct effect on occupational recruitment comes from father's occupation; whereas in the U.S. a larger effect comes from the respondent's education. The path from respondent's education to his previous occupation is not significant in the case of rural India, suggesting that education does not make much difference in the respondent's

initial placement in the occupational hierarchy. This however, was expected since we have already reported that there is very little variance in education of the respondents in this sample (Table 1). The absolute differences in terms of educational attainment of respondents in the two societies is hardly a new discovery. But what is interesting in this analysis is the finding that in both rural India and rural U.S. father's occupation has relatively little influence on respondent's education in both the raw metric comparisons of regression and in terms of path coefficients. The mean differences in respondent's education seem to differentiate rural sectors of industrial societies from rural sectors of agrarian societies.

There are two more features we want to add in this analysis. They are the examination of constants (α) and R^2 values of the equations. The R^2 values are uniformly higher for India (except for education, where the two are rather similar--having about zero values). Uniformly higher α values for the U.S. indicate the advantages American men have on these societal dimensions. The Indians seem to have only about two-thirds the advantage of U.S. men with regard to previous occupation. The effects of mean societal differences on the other status variables are too great to require explanations. They may be inspected (Table 2). What differences there are among different age cohorts remain to be seen.

Inter-Cohort Analysis of the Status Attainment Process

In this section we would like to examine the observed differences in the correlations among four age cohorts (Table 3). There is a monotonic increase in the magnitude of the correlation between a respondent's education and his current occupation in the rural United States, i.e., .302 for the oldest age

cohort and .404 for the youngest age cohort. Likewise, there is a monotonic increase in the magnitude of the correlation between previous occupation and current occupation in rural India. The correlation is .586 for the oldest age cohort and .927 for the youngest age cohort. There could be two interpretations for the rural U.S., i.e., (i) that the influence of education on occupational recruitment is increasing over time, or (ii) that the influence of education is greater at the initial entry into the job market and then declines. The interpretation with regard to the observed differences in the correlations in different age cohorts in India could be (a) that there is less occupational mobility for the younger age cohorts, and the chances of mobility increase as one becomes older (rather contrary to the trends observed in industrial countries and, for example, in our rural United States sample). Or alternatively (b) one can say that the chances of occupational mobility have declined over time. Inasmuch as no answers can be provided from the correlational analysis for these observed trends, we turn to the regression analysis of our data involving the same variables (Table 4).

Table 4 around here

In examining the regressors coefficients, we notice no serious distortions of the trends observed through correlational analysis. However, in the case of India, in all the age cohorts the regression coefficients from father's occupation to respondent's education are non-significant. Similarly regression coefficients leading from father's occupation to respondent's education, either to his previous occupation (i.e., the initial entry point into the labor force) or to his current occupation, are non-significant except in age cohort 45-54.

That coefficient, though significant, is not only too small to have any interpretation but also has a negative sign. When we observe raw regression coefficients for rural U.S., it becomes evident that the regression slope from education to current occupation does not change much from cohort to cohort, suggesting an almost uniform influence of education on current occupation over these age cohorts. For rural India, the regression slopes from previous occupation to respondent's current occupation remain similar to our correlational analysis. The magnitudes of both the standardized and unstandardized slopes show continued increase from older age cohorts to younger ones. This trend seems to have been reversed in the case of direct paths from father's occupation to respondent's current occupation.

Unstandardized regression coefficients, however, show a sharp contrast between rural Indian heads of households who entered the labor force before the 1920's and who experienced a fairly large direct effects of father's occupation on their own occupational attainment. This, however, changed dramatically for those who entered the labor force around 1923 and later. There may be an historical interpretation for this sharp increase in the direct effects of father's occupation, but one should also recognize that this increase does not occur until after these respondents are 45 years of age or older. The significance of this age becomes clear when one recognizes a unique feature of "return mobility" in rural India. By return mobility we mean the process by which sons return to the occupation of their fathers' after having spent a certain length of time in other occupations. Return mobility may occur because of economic pressures (early retirement, unemployment) or because of the demands of family values. For example, the eldest sons in rural India are expected to continue the family lines for ceremonial

purposes. The phenomenon of return mobility points to one important fact which is often ignored in the studies of social mobility, i.e., the fact of intergenerational mobility.

Return Mobility and the Life Cycle

A trend toward return mobility can be discovered from our rural India data. For example, an examination of Tables 5 and 6 indicates this tendency of clerical workers with farming origins to move back eventually to their father's occupation of agriculture (their destination status). Respondents with owner cultivator origins find their first entry and their destination to be the same to a large extent, but they are also the ones who enter into different occupations than their origins -- mostly into tenant farming, clerical, and sales. Similarly, though tenant farming receives entries from diverse occupational backgrounds, once a respondent becomes a tenant farmer, he tends to stay in that occupational category. Most of the new entry into tenant farming is by respondents whose fathers were owner cultivators or in sales positions. We might also point out that most of the respondents who went into sales occupations had fathers who were either in similar occupations or were tenant farmers. Several of the sales workers changed their positions to tenant farmers later in life. While other activity is interesting and, in general, points in the direction of our observations, we conclude that this phenomenon of "return mobility" in the life cycle process seems to explain the higher regression coefficients from father's occupation to son's current occupation for age cohorts 45-54 and 55-64 (Tables 4 and 5). Thus, "return mobility" seems a more plausible explanation than an alternative explanation that claims a dramatic increase in occupational inheritance around 1923 and after.

Predictability of the Life Cycle Processes

One way to study the predictability of life cycle processes such as those described here is by examining the R^2 values of different stratification models. The values of R^2 continue to increase in both samples as one goes from the older age cohorts to the younger age cohorts. This simply indicates that at the younger ages parental and other influences on status attainment are greater than when one gets older. However, the R^2 values for India are several times higher than those of the U.S., indicating considerably higher predictability in rural India than in rural U.S.

Conclusions

On the basis of these analyses, the following tentative conclusions are presented regarding similarities and differences between rural India and rural U.S. (Tables 3 and 4):

1.[a] Inheritance defined as $r^2_{fs(India)}=.676^2=.47$ $r^2_{fs\ (U.S.)}=.039^2=0$

2.[b] Direct Effects of Education: $Py_4v=.115$(India) $Py_4v=.266$ (U.S.)

3.[c] Direct Effects of Father's Occupation:

$Py_1x = .137$ to $Py_4x = .49$ (rises with age in India)

$Py_1x = .060$ to $Py_4x = -.012$ (falls with age, in anything, in the U.S.)

4. Total impact of all variables on Y_4:

Age	India	U.S.
25-34	.87	.25
35-44	.74	.15
45-54	.67	.15
55-64	.55	.11

5. As age increases in India, previous occupation has a diminishing effect, from large to moderate (.837, .673, .493, .370); as age increases in the U.S., first job status also has a diminishing effect, from moderate to small (.296, .124, .147, .148).

---

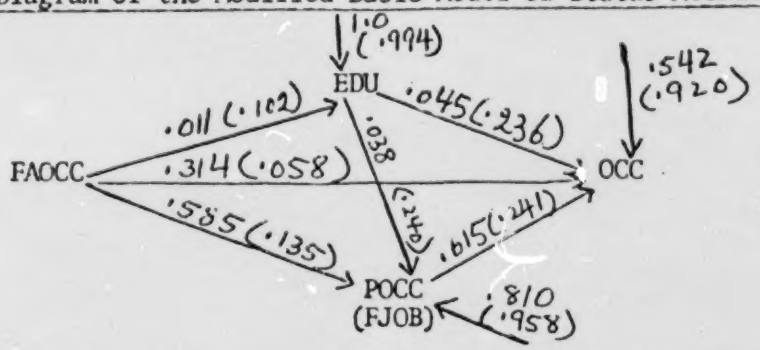
[a] f = father's occupational status at death or retirement in the case of rural India and at son's age 16 in the case of rural U.S. s = current occupational status for age cohort 55-64.

[b] y_4 = occupational status at age cohort 55-64. v = respondent's education.

[c] y = current occupational status (y_1 at age cohort 25-34; y_4 at age cohort 55-64). x = father's occupational status.

Figure 1

Path Diagram of the Modified Basic Model of Status Attainment



Note 1: Figures in parenthesis refer to the coefficients with regard to the rural United States (OCG-I data)

2: OCC = Respondents occupation in 1962
POCC = Previous occupation of the respondent
(FJOB) - (First Job of the respondent in the U.S.)
EDU = Education of the respondent
FAOCC = Father's Occupation

Figure 2

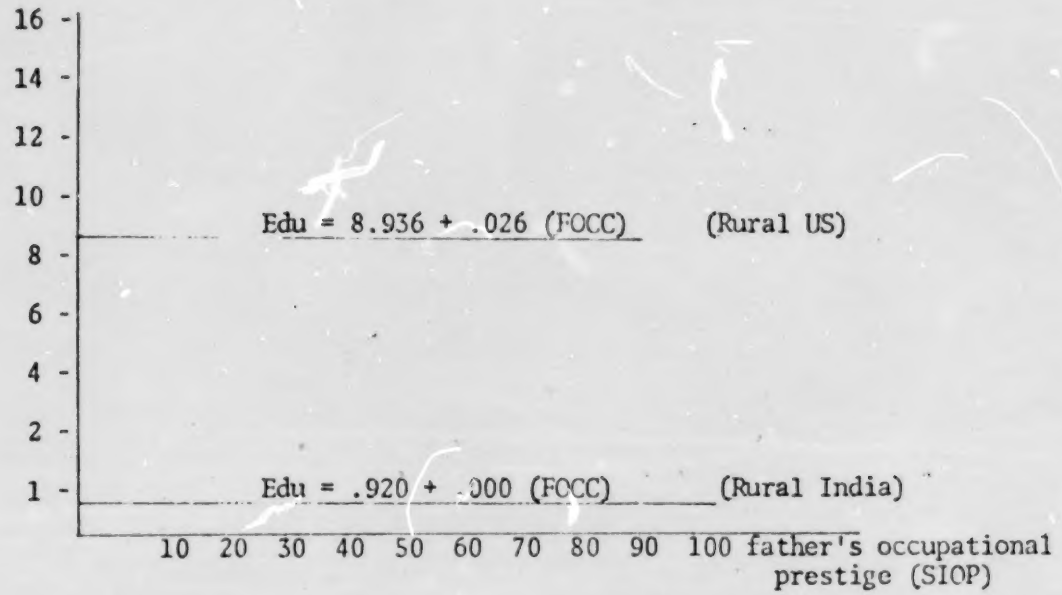

16
14
12
10
8
6
4
2
1
Edu = 8.936 + .026 (FOCC)
(Rural US)
Edu = .920 + .000 (FOCC)
(Rural India)
10 20 30 40 50 60 70 80 90 100 father's occupational prestige (SIOP)

TABLE 1

Coefficients of Correlation Between Selected Status Variables, Indian Male Heads of Households Aged 20-64/(Below diagonals) and United States Rural Males Aged 20-64/(Above diagonals), by Treimen's SIOP Scores.

	Prestige (SIOP) Metric					
	FAOCC	EDU	FJOB*	OCC	Mean	Standard Deviation
Father's Occupation (FAOCC)		.102	.156	.120	36.217	16.279
R's Education (EDU)	.001		.254	.303	10.000	3.584
R's Previous Occupation (POCC)	.585	.039		.310	31.515	12.823
R's Current Occupation (OCC)	.673	.069	.800		39.715	13.645
Mean	38.909	.931	37.672	37.987		
Standard Deviation	10.326	2.490	11.832	11.827		

* First job (FJOB) in the rural U.S. sample is compared with Respondent's previous occupation (POCC) in the rural Indian sample.

TABLE 2

Standard and Raw (Unstandardized) REgression Coefficients for Three Equation Stratification Model, Rural Indian Male Heads of Households and Rural United States Males, Aged 20-64 in 1962, by Treiman's SIOP Scale.

Dependent Variable	Independent Variables: Fathers's Occupation Ind.	Fathers's Occupation US	Education Ind.	Education US	Previous/First Occup. Job Ind.	Previous/First Occup. Job US	β Ind.	β US	R^2 Ind.	R^2 US
	STANDARDIZED COEFFICIENTS									
Education	.011	.102							.00	.01
Previous* Occupation (First Job)	.585	.135	.038	.240					.34	.08
Current Occupation	.314	.058	.045	.236	.615	.241			.70	.15
	RAW REGRESSION COEFFICIENTS									
Education	.000[a] (.009)	.026 (.003)					.920	8.936		
Previous Occupation (First Job)	.670 (.034)	.124 (.012)	.180 (.140)	.860 (.045)			11.425	18.000		
Current Occupation	.360 (.028)	.070 (.014)	.215 (.094)	1.072 (.056)	.614 (.024)	.306 (.016)	.647	13.09		

[a]Figures in parenthesis refer to standard error of estimates of raw regression.

* First job in the rural United States sample is compared with respondent's previous occupation in the rural Indian sample.

TABLE 3

Coefficients of Correlation between selected Status Variables for Rural Indian Male Heads of Households (above diagonal) and United States Rural Men (Below Diagonal) in 1962, for Four Age Cohorts.

Age Group and Variable	OCC	POCC	EDU	FAOCC	Mean Ind.	Mean US	Std. Dev. Ind.	Std. Dev. US	No of Respondents in the Sample Ind.	No of Respondents in the Sample US
25-34 (age 16 in 1943 to 1952)										
R's Occupation (OCC)		.927	.179	.658	38.81	36.53	12.79	15.73		
R's 1st Job (FJOB)	.412		.154	.617	36.18	32.18	13.06	13.54	178	1452
R's Education (EDU)	.404	.363		.106	1.16	11.07	2.88	3.36		
Father's Occupation (FAOCC)	.172	.208	.189		36.44	38.57	10.61	14.22		
35-44 (Age 16 in 1933 to 1942)										
R's Occuption (OCC)		.834	-.010	.673	38.01	38.47	11.33	14.93		
R's 1st Job (FJOB)	.233		.027	.606	37.77	32.00	11.33	12.13	186	1609
R's Education (EDU)	.361	.319		-.076	1.01	10.37	2.50	3.50		
Father's Occupation FAOCC)	.135	.128	.126		38.76	39.64	10.49	13.65		
45-54 (Age 16 in 1923-1932)										
R's Occupation (OCC)		.734	.035	.709	39.14	37.84	11.41	13.28		
R's Previous Occupation (POCC)	.241		-.016	.560	38.72	31.33	11.47	11.64	200	1337
R's Education (EDU)	.355	.271		.037	.97	9.20	2.53	3.56		
Father's Occupation (FAOCC)	.124	.145	.103		40.52	40.54	9.70	13.14		
54-65 (Age 16 in 1913-1922)										
R's Occupation (OCC)		.586	.005	.676	40.58	34.69	9.85	18.60		
R's Previous Occupation (POCC)	.214		-.275	.504	38.88	32.34	10.41	12.27	130	996
R's Education (EDU)	.302	.257		-.017	.48	8.22	1.74	3.57		
Father's Occupation (FAOCC)	.039	.180	.092		41.24	41.19	9.51	13.32		

TABLE 4

Standard and Raw (Unstandarized) Regression Coefficients for Three Equations Stratification Model, Rural Indian Male Heads of Households and Rural United States Males, and Four Age Cohorts, by Treiman's SIOP Scale

Dependent Variable and Age Cohort	Father's Occupation		Education		Previous First Occup. Job				R^2	
	Ind.	US	Ind.	US	Ind.	US	Ind.	US	Ind.	US
25-34 (Age 16 in 1943-52) STANDARDIZED COEFFICIENTS										
Education	.107	.189							.011	.04
Previous Occup. (First Job)	.608	.145	.089	.335					.38	.15
Current occup.	.137	.060	.035	.286	.837	.296			.87	.25
35-44 (Age 16 in 1933-1942)										
Education	-.076	.126							.005	.02
Previous occup. (First Job)	.611	.089	.073	.307					.37	.11
Current occup.	.265	.080	-.008	.311	.673	.124			.74	.15
45-54 (Age 15 in 1923-32)										
Education	.037	.103							.001	.01
Previous occup. (First Job)	.561	.118	-.037	.259					.31	.09
Current occup.	.433	.071	.027	.308	.493	.147			.67	.15
55-65 (Age 15 in 1913-1922)										
Education	-.015	.092							.017	.01
Previous occup. (First Job)	.499	.158	-.267	.242					.32	.09
Current occup.	.492	-.012	.115	.266	.370	.148			.55	.11
25-34 RAW REGRESSION COEFFICIENTS										
Education	.029	.045					.100	9.340		
	(.020)	(.006)								
Previous occup. (First Job)	.748	.138	.403	1.350			8.541	11.929		
	(.073)	(.023)	(.269)	(.099)						
Current occup.	.166	.062	.156	1.335	.820	.344	-.063	8.300		
	(.047)	(.026)	(.121)	(.115)	(.034)	(.029)				

Table 4 (con't)

	Father's Occupation		Education		Previous First Occup.		Job		R^2	
	Ind.	US	Ind.	US	Ind.	US	Ind.	US	Ind.	US
35-44										
Education	-.018	.032					1.706	9.091		
	(.017)	(.006)								
Prev. Occup. (First Job)	.660	.079	.331	1.066			11.861	17.803		
	(.063)	(.021)	(.266)	(.082)						
Current Occup.	.295	.088	-.036	1.328	.694	.153	.398	16.319		
	(.053)	(.026)	(.178)		(.048)	(.030)				
45-54										
Education	.010	.028					.583	8.061		
	(.018)	(.007)								
Prev. Occup. (First Job)	.663	.105	-.167	.848			12.010	19.270		
	(.070)	(.023)	(.028)	(.086)						
Current Occup.	.509	.082	.123	1.323	.490	.194	-.553	16.274		
55-64										
Education	-.003	.025							.609	7.211
	(.016)	(.008)								
Prev. Occup. (First Job)	.547	.145	-1.597	.831					17.107	19.512
	(.080)	(.028)	(.436)	(.104)						
Current Occup.	.509	-.016	.648	1.380	.350	.224			5.675	16.765
	(.072)	(.042)	(.356)	(.161)	(.069)	(.048)				

[a] Standard error of estimates are in the parenthesis.

Table 5

Occupational Changes from Father's Occupation (ORIGIN) to Previous Occupation (ENTRY) Rural Indian Male Heads of Households, Ages 20-64, in 1962, by Broad Occupational Classifications

Father's Occupation (Origin)

Respondent's previous occupation (entry)	Professional, Administrators	Clerical Workers	Owner Cultivator	Sales Worker	Tenant Farmers	Skilled & Semi-Skilled Workers	Service Workers	Agricultural Workers	Total
Hi	1(3)	2(5)	3(1)	4(4)	5(2)	6(6)	7(8)	8(7)	
1	3(33.3)	0(00.0)	4(0.10)	2(2.3)	1(1.2)	3(4.3)	0(00.0)	0(00.0)	13
2	0(00.0)	5(62.5)	22((05.3)	4(4.5)	3(3.6)	1(1.4)	3(6.1)	2(5.6)	40
3	1(11.1)	1(12.5)	328(78.8)	4(4.5)	23(28.0)	5(7.1)	3(6.1)	0(00.0)	365
4	4(44.4)	0(00.0)	10(02.4)	35(39.8)	3(3.6)	0(00.0)	0(00.0)	3(8.3)	55
5	0(00.0)	0(00.0)	30(07.2)	32(36.4)	35(42.7)	3(4.3)	3(6.1)	0(00.0)	103
6	1(11.1)	1(12.5)	8(01.9)	1(1.1)	6(7.3)	44(62.8)	2(4.1)	0(00.0)	63
7	0(00.0)	1(12.5)	10(02.4)	9(10.2)	9(11.0)	9(12.8)	37(75.5)	7(19.4)	82
8	0(00.0)	0(00.0)	4(01.0)	1(1.1)	2(2.4)	5(7.1)	1(2.0)	24(66.7)	37
Total	9(100.0)	8(100.0)	416(100.0)	88(100.0)	82(100.0)	70(100.0)	49(100.0)	36(100.0)	758

Hi — Lo

Note: (1) Missing observations were not included in the analysis.
(2) Numbers from high to low refers to the ranking of broad occupational classifications by SIOPS measures. Numbers in parentheses are percentages.

-77-

Table 6

Occupational Changes from Previous Occupation (Entry) to Current Occupation (Destination) Rural Indian Male Heads of Households, Aged 20-64 in 1962, by Borad Occupational Classification

Current Occupation (Destination)	Previous Occupation (ENTRY)								
	Professional, Administrators	Clerical Workers	Owner Cultivators	Sales Workers	Tenant Cultiv.	Skilled & Semi-Skilled Workers	Service Workers	Agricultural Laborers	Total
1	10(76.9)	2(5.1)	0	1(1.8)	0	1(1.6)	0	0	13
2	0	15(38.5)	3(0.85)	0	0	0	0	0	18
3	3(23.1)	13(33.3)	339(92.6)	5(8.9)	0	4(8.5)	0	0	371
4	0	2(5.1)	4(1.1)	28(50.0)	0	0	0	0	34
5	0	3(7.7)	3(.08)	19(33.9)	103(98.1)	3(4.8)	1(2.7)	1(2.7)	138
6	0	2(5.1)	6(1.6)	0	0	52(82.5)	1(2.7)	1(2.7)	64
7	0	2(5.1)	10(2.7)	2(3.6)	0	2(3.2)	0	0	82
8	0	1(2.6)	1(0.3)	1(1.8)	2(1.9)	1(.16)	35(94.4)	35(94.4)	49
Total	13(100.0)	39(100.0)	366(100.0)	56(100.0)	105(100.0)	63(100.0)	37(100.0)	83(100.0)	761

Note: (1) Missing observations were not included in the analysis.
(2) Numbers from high to low refer to the ranking of broad occupational classifications by SIOPS measures. Numbers in parentheses are percentages.

FOOTNOTES

1. 'Paradigm' has been defined by Thomas Kuhn as "a term that related closely to normal science." By normal science Kuhn means "to suggest that some accepted examples of actual scientific practice--examples which include law, theory, application, and instrumentation together provide models from which spring particular coherent traditions of scientific research." (Kuhn, 1962:10)

2. This reconceptualization has been the result of several conceptual and methodological problems in the analysis of social mobility. See, for example, Duncan's critique of the traditional methods of analyzing mobility data (Duncan, 1966). Blau and Duncan's paradigm is sometimes considered to have brought a "breakthrough" (defined as a shift in the dominant paradigm) in mobility research.

3. For a discussion and definition of "analytic indicator," see Sheldon and Land (1972). Featherman, et al. (1974) argue that Blau and Duncan's paradigm of stratification research fits very well into Sheldon and Land's definition of an analytic indicator. In order to provide the justification that the process of stratification is of interest to other scholars, they point out that "Land (1971) has argued that from the perspectives of both the student of society and the monitor-reporter of social trends, social indicators will become most valuable (in units of information) when they are chosen from the framework of theory and conceptions of how society and its institutions work; they should be components of social process models." (Featherman, et al., 1974 b:19)

4. We are thankful to Professor David L. Featherman for his permission to use these data for this research.

5. We are grateful to Dr. B. K. Roy Burman, Deputy Registrar, Census of India (Dr. Roy Burman heads the Social Study Division within the Census organization) for providing us with the unpublished data. For a complete description of this data set, see Elder and Sharda (in progress).

6. In criticizing the 'disappearance of rural subculture hypotheses in the U.S. by economists' (see for example Boulding, 1963); rural-urban continum hypothesis of anthropologists and sociologists (for example Queen, and Carpenter, 1953), etc., Schnore points out that the "place of residence is a fundamental social characteristic that merits the same attention given such attributes as age, sex, ethnicity, and socioeconomic status. This cluster of traits is the most important set of individual characteristics with which sociologists deal. Indeed, it is rather difficult to add to this list of statuses without leaving the sociological domain altogether and beginning to characterize people according to purely psychological traits." "Not only is place of residence--rural versus urban--a crucial current characteristic; it is also a vital variable when regarded from the standpoint of one's place of origin. A very wide range of behavior can be predicted with reference to either (a) the type of community in which the person now resides, or (b) the type of community in which he was born and reared." (Schnore, 1966:136)

7. Treiman's STANDARD INTERNATIONAL OCCUPATIONAL PRESTIGE SCALE (SIOPS) is correlated about .7 (high) both with Duncan's SEI (1962) in the U.S. (Featherman, et al., 1974 a:Table 6) and with Sharda's Rural Indian Occupational Prestige Scale (Sharda, 1975:Table 5), indicating high common variance in all the three scales when employed in mobility studies. Treiman's scale, however, is the most comprehensive one and hence most suitable for intersocietal comparison. It may be pointed out here that Treiman's SIOPS is a prestige scale; whereas the other two scales are socio-economic scales of their societies. However, the estimation of social mobility of individuals through this prestige scale is probably not in error by more than a small 'arithmatic factor' (because of SIOP's high correlation with the local scale) in any statistical estimation of cause and effect. For similar arguments see Featherman, et al. (1974 a:14).

8. Path analysis as a method for analyzing social processes was introduced to sociology by Duncan (1966). For further refinements see Land (1969), Finney (1972), and Duncan, Featherman, and Duncan (1972:chap. 2).

BIBLIOGRAPHY

Blau, Peter M., and Otis D. Duncan
1967 The American Occupational Structure, New York: John Wiley and Sons

Boulding, Kenneth E.
1963 "The Death of the City: A Frightened Look at Post Civilization," in Oscar Handlin and John Burchard (eds.), The Historian and the City, Cambridge: M.I.T. Press

Cutright, Phillips
1968 "Occupational Inheritance: A Cross-National Analysis" AJS, 73 (January): 400-416

Duncan, Otis D.
1961 "A Socio-economic for all Occupations," pp. 109-138 in Albert J. Reiss, et al., (eds.); Occupations and Social Status, New York: Free Press of Glencoe

Duncan, Otis D.
1966 "Methodological Issues in the Analysis of Social Mobility," pp. 51-97 in Neil J. Smelser and Seymour M. Lipset (eds.), Social Structure and Mobility in Economic Development, Chicago: Alnine

Duncan, Otis D.
1966 "Path Analysis: Sociological Examples" AJS, 72 (July): 1-16

Duncan, Otis D., David L. Featherman, and Beverly Duncan
1972 Socioeconomic Background and Achievement, New York: Seminar Press

Elder, Joseph W., and Bam Dev Sharda
In
Progress Census of India Village Survey Monograph Project

Featherman, David L., William H. Sewell, and Robert M. Hauser
1974a "Comparable Data on Inequality and Stratification: Perspectives on the Second Generation of National Mobility Studies," The American Sociologist, 9 (February): 18-25

Featherman, David L., F. Lancaster Jones, and Robert M. Hauser
1974b "Assumptions of Social Mobility Research in the U.S.: The Case of Occupational Status," Working Paper 74-77, Centre for Demography and Ecology, Madison: University of Wisconsin

Finney, John M.
1973 "Indirect Effects in Path Analysis," Sociological Methods and Research, 1 (November): 175-186.

Iutaka, Sugiyama and E. Wibur Bock
1973 "Determinants of Occupational Status in Brazil," pp. 213-222 in Walter Müller and Karl Ulrich Mayer (eds.), Social Stratification and Career Mobility, Paris: Mouton

Kuhn, Thomas S.
1962 The Structure of Scientific Revolutions, Chicago: University of Chicago Press

Safar, Adenek
1971 "Different Approaches to the Measurement of Social Differentiation of the Czechoslovak Socialist Society," Quality and Quantity, 5(1): 179-208

Schnore, Leo F.
1966 "The Rural-Urban Variable: An Urbanite's Perspective," Rural Sociology, 31(June): 131-143

Sheldon, E. B., and Kenneth C. Land
1972 "Social Reporting for the 1970's," Policy Sciences, 3:137-151

Sharda, Bam Dev
1974 Status Attainment in Rural India, unpublished Ph.D. dissertation, University of Wisconsin, Madison.

Sharda, Bam Dev
1975 "Occupational Prestige: Scales for Rural India and Their Bases," paper to be presented at the 70th Annual Meeting of the American Sociological Association, San Francisco, August 25-29, 1974

Land, Kenneth C.
1969 "Principles of Path Analysis," in Edgar F. Borgatta and George W. Bohrnstedt (eds.), Sociological Methodology, San Francisco: Jossey-Bass Inc.

Land, Kenneth C.
1971 "On the Definition of Social Indicators," The American Sociologist, 6(November): 322-325

Lenski, Gerhard E.
1966 Power and Privilege: A Theory of Social Stratification, New York: McGraw-Hill

Treiman, Donald J.
1970 "Industrialization and Social Stratification," in Edward O. Laumann (ed.), pp. 207-234 in Social Stratification: Research and Theory for the 1970s, Indianapolis: Bobbs Merrill

Treiman, Donald J.
Forthcoming
"Standard International Occupational Prestige Scale," New York: Academic Press.